

# Gaussian Process Models for Mortality Rates and Improvement Factors: An Interactive R Markdown Approach

AMS Sectional Meeting AMS Special Session on Markov Processes, Gaussian Processes and Applications

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## Mortality Rates

- Consider some national mortality data
- Observe  $D(x_{ag}, x_{yr})$  **deaths** for an individual aged  $x_{ag}$  in calendar year  $x_{yr}$
- Out of  $E(x_{ag}, x_{yr})$  **exposures**
  - ▶ e.g. 100,000 individuals aged 50 in 2008, and 1,000 of those individuals die throughout the year, then

$$D(50, 2008) = 1000, \quad E(50, 2008) = 100,000$$

- **Central mortality rate** is defined as

$$e^{-\mu(x_{ag}, x_{yr})} = \frac{D(x_{ag}, x_{yr})}{E(x_{ag}, x_{yr})} = \frac{1000}{100000} = 0.01$$

- Mortality models model the **log-mortality rate**

$$\mu(x_{ag}, x_{yr})$$

which is roughly **linear** in age (increasing) and calendar year (decreasing)

# Mortality Modeling

- (~1850) Gompertz-Makeham Law

$$\mu(x_{ag}, x_{yr}) \equiv \mu(x_{ag}) = ax_{ag} + b \exp(cx_{ag} + d)$$

- ▶ Doesn't take into account calendar year (technological advances)

- (1992) Lee-Carter

$$\mu(x_{ag}, x_{yr}) = \alpha_{x_{ag}} + \beta_{x_{ag}} \kappa_{x_{yr}}$$

where  $\alpha, \beta, \kappa$  are separately modeled as time series,  
e.g.  $\kappa_{x_{yr}} \sim \text{ARIMA}(0, 1, 0) + \text{drift}$

- ▶ Began a trend of **stochastic** mortality models
- ▶ Captures more complicated dynamics
- ▶ Allows for uncertainty quantification
- ▶ Highly parametric

# Mortality Modeling

- Extensions of Lee-Carter
  - ▶ Adding cohort effect  $+\gamma_{x_{yr}-x_{ag}}$
- Spline Models
  - ▶ Nonparametric
  - ▶ Hard to quantify uncertainty
- (2018) Ludkovski-Risk-Zail

$$\mu(x_{ag}, x_{yr}) \sim GP(m, C)$$

- ▶ Models as a Gaussian process with prior mean  $m$  and covariance kernel  $C$
- ▶ Nonparametric
- ▶ Allows for uncertainty quantification
- ▶ Bayesian (easy to update)
- ▶ Handles edge uncertainty
- ▶ Handles missing data (tough with time series)

# Mortality Improvement

- **Mortality Improvement** involves analyzing

$$MI(x_{ag}, x_{yr}) = \mu(x_{ag}, x_{yr}) - \mu(x_{ag}, x_{yr} - 1),$$

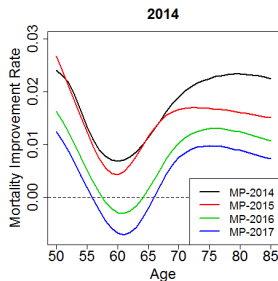
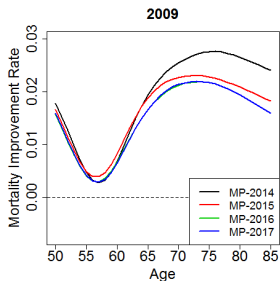
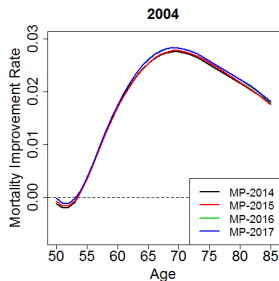
how mortality improves **over time**.

- Typically the **one-year difference** is analyzed (as above)
- We expect  $MI > 0$  due to medical advances
- The Society of Actuaries (SOA) publishes a study  $MP$  every year to analyze **US** mortality improvement

# US Mortality Improvement

## MP-2014 / MP-2015

- Published by SOA, uses “RPEC\_2014” model
- US CDC Data
- MP-2014 uses years 1950-2009
- Plans to update scales at least triennially; two years of additional CDC data shows drastic change in later years
  - ▶ MP-2015 emerges
  - ▶ (Also showing MP-2016, MP-2017, came after this paper finished)



## Goal:

Model US Mortality data using Gaussian Process (GP) regression

- Bayesian
- Provides posterior Gaussian distribution for input of **any** age and year
- Offers easy analysis of both mortality and mortality improvement simultaneously
- Gaussian distribution implies one-year mortality improvement factors remains Gaussian
- Differentiable: can provide **instantaneous** mortality improvement (still Gaussian)
- Spatial approach inherently handles missing and edge data
- Provide simple to use code with output through an **R notebook**

# Typical Regression Assumption

Hypothesis:

$$y = f(x) + \varepsilon$$

- **Observe**  $\mathbf{y} = y^{1:N}$  for **input locations**  $\mathbf{x} = x^{1:N}$
- Want to understand the **function**  $f$ 
  - ▶ e.g.  $f(x) = \beta_0 + \beta_1 x$  (simple linear regression)
- $\varepsilon$  is **noise**
  - ▶ e.g. measurement error
  - ▶ can't observe  $f(x)$  directly
- Assume  $\varepsilon \sim \mathcal{N}(0, \sigma^2(x))$  (often  $\sigma(x) \equiv \sigma \in \mathbb{R}^+$ )
- Often in mortality modeling:  $f(x)$  is based on an **ARIMA** process(es) or on **splines**
- Our assumption:  $f$  is a Gaussian Process (modeling log-mortality,  $x = (x_{ag}, x_{yr})$ )



# Gaussian Process

- Defined as a collection of random variables  $\{f(x) | x \in \mathbb{R}^d\}$
- Any finite subset has a multivariate Gaussian distribution with covariance  $C(\cdot, \cdot)$ :

$$f(x_1), \dots, f(x_n) \sim \mathcal{N} \left( (m(x_1), \dots, m(x_n)), C(\mathbf{x}, \mathbf{x}^T) \right).$$

- Fix mean function  $m$  and covariance kernel  $C$ ; this provides a prior distribution

# Modeling with Gaussian Processes

- 1 Declare prior **mean** function and covariance kernel
  - ▶ Mean function can also be parametric and fitted with data; useful in extrapolation
  - ▶ Covariance kernel governs spatial relation between points
  - ▶ Hyperparameters can be specified using expert knowledge or fitted from data

# Modeling with Gaussian Processes

- 1 Declare prior **mean** function and covariance kernel
  - ▶ Mean function can also be parametric and fitted with data; useful in extrapolation
  - ▶ Covariance kernel governs spatial relation between points
  - ▶ Hyperparameters can be specified using expert knowledge or fitted from data
- 2 Output can be easily evaluated at **any** location
  - ▶ Output is a **random variable** with mean and covariance depending on neighboring inputs

## Posterior

- Observe **pairs**  $(\mathbf{y}, \mathbf{x}) = ((y, x)^{1:N})$ 
  - ▶ (e.g.  $y$  = historic log-mortality and  $x$  = (age, year))
- Gaussian assumptions imply that marginally for any input  $x$

$$f(x)|(\mathbf{y}, \mathbf{x}) \sim \mathcal{N}\left(m_*(x), s_*^2(x)\right)$$

## Posterior

- Observe **pairs**  $(\mathbf{y}, \mathbf{x}) = ((y, x)^{1:N})$ 
  - (e.g.  $y =$  historic log-mortality and  $x =$  (age, year))
- Gaussian assumptions imply that marginally for any input  $x$

$$f(x)|(\mathbf{y}, \mathbf{x}) \sim \mathcal{N}\left(m_*(x), s_*^2(x)\right)$$

- $m_*$  and  $s_*^2$  are the posterior mean and variance functions

$$\begin{cases} m_*(x) \doteq \mathbf{c}(x)^T (\mathbf{C} + \Sigma)^{-1} \mathbf{y}; \\ s_*^2(x) \doteq C(x, x) - \mathbf{c}(x)^T (\mathbf{C} + \Sigma)^{-1} \mathbf{c}(x), \end{cases} \quad (1)$$

where

$$\begin{cases} \mathbf{c}(x) \doteq \left( C(x, x^i) \right)_{1 \leq i \leq N} \quad (\text{covariances between } x \text{ and inputs } \mathbf{x}) \\ \mathbf{C} \doteq \left( C(x^i, x^j) \right)_{1 \leq i, j \leq N} \quad (\text{covariances between inputs } \mathbf{x}) \\ \Sigma \doteq \text{diag} \left( \sigma^2(x^i) \right) \quad (\text{diagonal matrix of noise variance}) \end{cases}$$

## Covariance Kernels & Parameter Estimation

- Common choice is squared-exponential (or Gaussian) covariance kernel

$$C(x, x') = \eta^2 \exp \left( -\frac{(x_{ag} - x'_{ag})^2}{2\theta_{ag}^2} - \frac{(x_{yr} - x'_{yr})^2}{2\theta_{yr}^2} \right).$$

- Knowing mortality at  $x$  will greatly influence mortality at "neighboring"  $x$ 's
  - e.g. knowing mortality for a **80 year old in 2015** greatly aids in prediction of a **85 year old's mortality in 2016**; knowing a **50 year old's mortality in 2000** has a nearly non-existent effect
- Implies **hyperparameter** family of  $\Theta \doteq (\theta_{ag}, \theta_{yr}, \eta^2, \sigma^2)$ 
  - Also mean function hyperparameters (if included)
- Estimates** are fit using MLE; likelihood can be written out explicitly due to Gaussian assumptions
  - Done using R package `DiceKriging`
- Alternatively, can use Bayesian approach with priors on  $\Theta$ 
  - Separate package using STAN language
  - Leads to non-Gaussian posterior

## Illustrative Example

Goal: Learn  $f(x) = \sin(x)$  over domain  $[0, 1.5\pi]$

- Observe realizations of

$$y = \sin(x) + \varepsilon$$

where  $\varepsilon \sim N(0, \sigma = 0.5)$

- Try:
  - ▶  $x = 0.25, 0.5, 0.75, \dots, 2.75, 4.5$  ( $N = 18$ )
- Then **update** model with data on  $(1.5\pi, 2\pi]$  to see how the overall fit changes
  - ▶ add  $x = 4.75, \dots, 6$  (total  $N = 24$ )

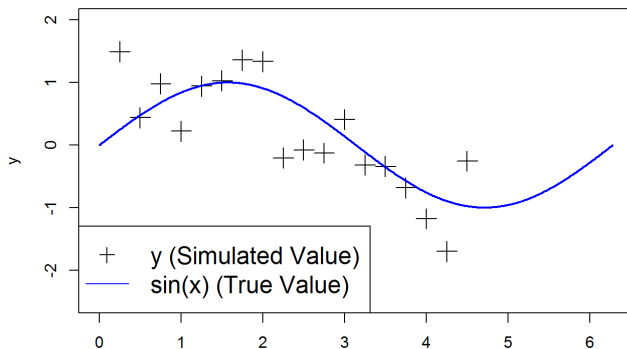
# Illustrative Example

$$y = \sin(x) + \varepsilon, \quad \varepsilon \sim \mathcal{N}(0, 0.01x)$$

- 1 Generate data from random process

$$y = \sin(x) + \varepsilon, \quad e \sim N(0, 0.5)$$

18 Design Points, [0,4.5]

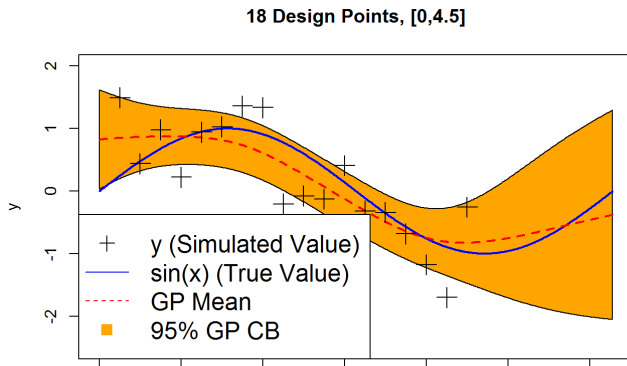




## Illustrative Example

Fit GP to  $N = 18$   $(x, y)$  pairs

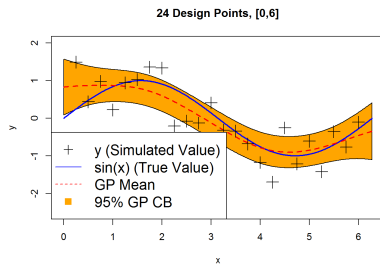
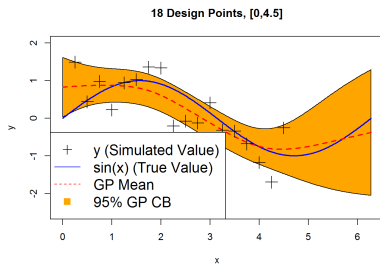
- Estimate hyperparameters  $(\theta, \eta, \sigma)$
- Produce posterior mean, covariance matrix (provides credible intervals)
- Observe **naturally** increasing uncertainty at edge data



# Illustrative Example

Update GP with  $N_{new} = 6$   $(x, y)$  pairs on  $[4.75, 6]$

- (Optional) update hyperparameters  $(\theta, \eta, \sigma)$
- Produce posterior mean, covariance matrix (provides credible intervals)



## R Notebook Code

```
x <- seq(0.25,4.5,0.25)
n <- length(x)
y <- sin(x)+rnorm(n,0,0.5)
```

```
library(DiceKriging)
fit_nug <- km(formula = ~1,
              design = data.frame(x=x), response = y,
              nugget.estim = TRUE,
              covtype= "gauss",
              optim.method="gen")
```

```
nug <- fit_nug@covariance@nugget
```

```
fit <- km(formula = ~1,
          design = fit_nug@X, fit_nug@y,
          noise.var = rep(nug, fit_nug@n),
          coef.trend = fit_nug@trend.coef, #re-use obtained hyperparameters
          coef.cov = fit_nug@covariance@range.val)
```

```
##
## Call:
## km(formula = ~1, design = fit_nug@X, response = fit_nug@y, coef.trend = fit_nug@trend.coef,
##   coef.cov = fit_nug@covariance@range.val, noise.var = rep(nug,
##   fit_nug@n))
##
## Trend coeff.:
##
## (Intercept)  -0.0346
##
## Covar. type : matern5_2
## Covar. coeff.:
##
##   theta(x)   1.3373
##
## Variance: 0.5803238
```

# Comments

- $m(x) = m$  assumed (*clearly not true*)
- In practice,
  - ▶ Data is usually detrended, or
  - ▶ Parametric trend function e.g.  $f(x) = \beta_0 + \beta_1 x$  assumed (and fitted alongside)
- Example is **one-dimensional** ( $x \in \mathbb{R}$ )
  - ▶ Framework naturally extends to multi-dimensional case ( $x \in \mathbb{R}^d$ ), for example
    - ★  $f(x, y) = \sin(x) \cos(y) + 2xy$
    - ★  $f(\text{age}, \text{year}) = (\text{mortality rate depending on age, year})$

# R Notebook For GPs on Mortality Data

- Provide R “code blocks” along with explanation of what it does and discussion of results
- Practitioner can choose to modify as much as needed
  - ▶ For example, simply change `cdcMale.csv` to `myInsuranceCompanyData.csv`
  - ▶ Plots have changeable ranges (easy to choose what years to plot)
  - ▶ **ALL** code is available, so programmers can easily modify as needed

# Data

## CDC Data

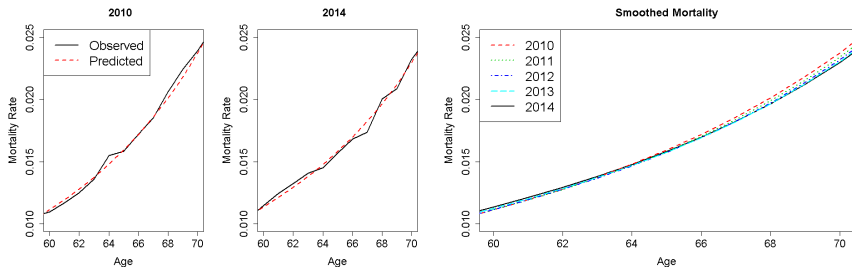
- United States
- Ages 50–84, Years 1999–2014
  - ▶ 1360 Data Points ( $x = (x_{ag}, x_{yr})$ )
  - ▶ 84 is maximal age for CDC data
  - ▶ 50 chosen as cutoff to minimize mixing lower age behavior
  - ▶ 1999 earliest year available on wonder.cdc.gov
  - ▶ Could add earlier years, but our analysis suggests they have little effect
  - ▶ Most relevant for **longevity risk**

# GP Model Assumptions

- Observe **central mortality rate**  $e^{-\mu(x_{ag}, x_{yr})} = D(x_{ag}, yr) / E(x_{ag}, yr)$
- Fit log-mortality rate  $y$  to  $x = (x_{ag}, x_{yr})$  pairs
- Can try  $\sigma(x)$  based on Binomial assumption
  - ▶ Overdispersion issues ( $\mu_{ag, yr}$  is unknown)
  - ▶ Minimal change in final model from simply choosing  $\sigma := \sigma(x)$
- Use Gaussian covariance kernel
  - ▶ Implies  $f$  is differentiable
  - ▶ Minimal change in final model from other kernel choices
- Changed  $m(x)$  trend function based on application
  - ▶ In-sample analysis generally used  $m(x) = \beta_0$
  - ▶ Out-of-sample generally used  $m(x) = \beta_0 + \beta_1 x_{ag} + \beta_2 x_{ag}^2 + \beta_3 x_{yr}$  (like Gompertz)

# Posterior Predicted Mortality Rates

- Showing  $m_*(x)$  for each ages 60–70
- Left panels include historic observations
- Right panel suggests mortality improvement





# Goals

- In-sample smoothing
- Extrapolation (both in calendar year and age)
- Mortality Improvement

$$MI_{back}^{obs}(x_{ag}; yr) \doteq 1 - \frac{\exp(\mu(x_{ag}, yr))}{\exp(\mu(x_{ag}, yr - 1))}$$

compare with SOA MP-2015 results

## Loading Data

- Input data should be an R data frame with
  - ▶ age, calendar year, deaths, exposure
- The corresponding log mortality rates are computed as

$$y^n = \log(D^n/L^n)$$

- ▶  $D^n$  is the number of deaths and for the  $n$ th age/year pair  
 $x^n = (x_{ag}^n, x_{yr}^n)$ ,
- ▶  $L^n$  midyear count of lives

```
mortData <- read.csv("cdcMale.csv",header=T)
mortData$rate <- mortData$D / mortData$L
mortData$y <- log(mortData$rate)
head(mortData)
```

```
##   X age year      D      L      rate      y
## 1 1  50 1999  9775 1847555 0.005290776 -5.241790
## 2 2  51 1999 10470 1762492 0.005940452 -5.125970
## 3 3  52 1999 11509 1900702 0.006055131 -5.106849
## 4 4  53 1999  9885 1355175 0.007294261 -4.920667
## 5 5  54 1999 10717 1413117 0.007583944 -4.881722
## 6 6  55 1999 11728 1390616 0.008433673 -4.775523
```

# Fitting the Model

```
xMort <- data.frame(age = mortData$age, year = mortData$year)
yMort <- mortData$y
mortModel_nug <- km(formula = ~1,
                    design = data.frame(x = xMort), response = yMort,
                    nugget.estim=TRUE,
                    covtype="gauss",
                    optim.method="gen",
                    # the "control" parameters below handle speed versus risk of
                    # converging to local minima. See "rgenoud" package for details
                    control=list(max.generations=100,pop.size=100,wait.generations=8,
                                solution.tolerance=1e-5))

mortModel_nug

##
## Call:
## km(formula = ~1, design = data.frame(x = xMort), response = yMort,
##     covtype = "gauss", nugget.estim = TRUE, optim.method = "gen",
##     control = list(max.generations = 100, pop.size = 100, wait.generations = 8,
##                   solution.tolerance = 1e-05))
##
## Trend  coeff.:
##              Estimate
## (Intercept) -3.8710
##
## Covar. type : gauss
## Covar. coeff.:
##              Estimate
## theta(x.age)  15.8250
## theta(x.year) 15.5361
##
## Variance estimate: 1.842994
##
## Nugget effect estimate: 0.0002808436
```

# In-Sample Smoothing

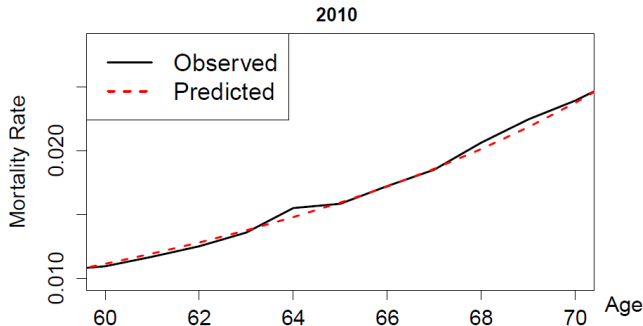
```

agesForecast <- seq(58,72,1)
yearsForecast <- 2010

rateForecastTrue <- dplyr::filter(mortData,age %in% agesForecast, year %in% yearsForecast)$rate
# build data frame for desired forecasts, then call predict
xPred <- data.frame(age=agesForecast,year=yearsForecast)
mortPred <- predict(mortModel, newdata=data.frame(x=xPred),
                   cov.compute=TRUE,
                   se.compute=TRUE,type="UK")

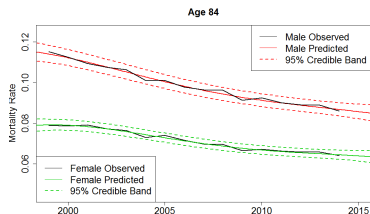
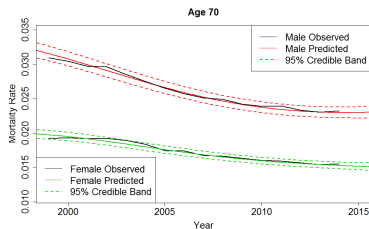
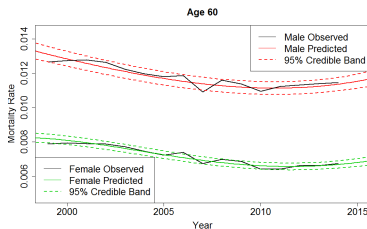
# plot mortality as a function of age
plot(agesForecast,rateForecastTrue,type="l",lwd=2, main="2010",
     xlab="Age",ylab="Mortality Rate",cex.axis=1.3,cex.lab=1.3,cex=1.3,
     xlim=c(60,70))
lines(agesForecast,exp(mortPred$mean),col=2, lty=2,lwd=2)
legend("topleft",c("Observed","Predicted"),lwd=c(2,2),lty=c(1,2),col=c(1,2),cex=1.5)

```



# Mortality Over Time with Credible Bands

- Posterior mean and 95% credibility bands for  $f_*$  over calendar year
- Can observe increasing **uncertainty at edges**
- Observe mortality improvement then **decline**

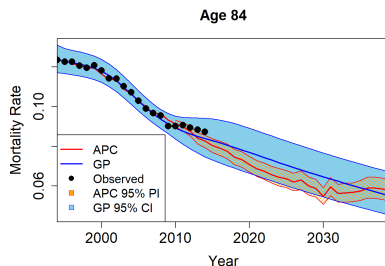
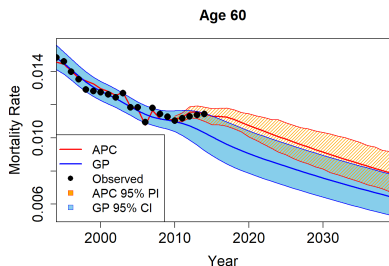


# Extrapolation

- Compare to `apc` model from R package `stmomo`

$$\mu(x, t) = \alpha(x) + \kappa(t) + \gamma(t - x), \quad \sum_c \gamma(c) = 0, \quad \sum_c c\gamma(c) = 0$$

- GP produces similar forecasts with more desirable smoothness properties



# Mortality Improvement

- Typical way is to look at the **annual backward improvement**

$$MI_{back}^{obs}(x_{ag}; yr) \doteq 1 - \frac{\exp(\mu(x_{ag}, yr))}{\exp(\mu(x_{ag}, yr - 1))}$$

- $f_*(x_{ag}, yr)$  is a random variable, so we have the **predicted mean improvement**

$$m_{back}^{GP}(x_{ag}, yr) = \mathbb{E} \left[ MI_{back}^{GP}(x_{ag}, yr) \right] \doteq \mathbb{E} \left[ 1 - \frac{\exp(f_*(x_{ag}, yr))}{\exp(f_*(x_{ag}, yr - 1))} \right]$$

- ▶ Available in closed form (lognormal distribution)
- Also have  $MI_{back}^{MP}(x_{ag}; yr)$  (published MP-2015 improvement factors)

# YoY Mortality Improvement Plots

```

# compare to MP-2015 tables
mp2015 <- read.csv("mp2015.csv")

agesForecast <- 48:86
agesObserved <- 50:84
yearsForecast <- c(2000,2014)
yearsPred <- c(1999,2000,2013,2014) # need to add one year back for improvement

nYr <- length(yearsPred)
nAg <- length(agesForecast)

# predict
xPred <- data.frame(age=rep(agesForecast,each=nYr),year=rep(yearsPred,nAg))
mortPred <- predict(mortModel, newdata=data.frame(x=xPred),cov.compute=TRUE,
  se.compute=TRUE,type="UK")

xPred$m <- mortPred$mean
for(yr in yearsForecast){
  forwardObs <- dplyr::filter(mortData, age %in% agesObserved, year == yr)$rate
  backwardObs <- dplyr::filter(mortData, age %in% agesObserved, year == yr-1)$rate
  M1backobs <- 1-forwardObs/backwardObs # raw observed improvement rates

  forwardPred <- filter(xPred, age %in% agesForecast, year == yr)$m
  backwardPred <- filter(xPred, age %in% agesForecast, year == yr-1)$m
  # smoothed improvement rates using the GP model
  m1backgp <- 1-exp(forwardPred)/exp(backwardPred)

  plot(agesObserved,M1backobs, type="l", lwd=2, main = yr, ylim=c(-0.05, 0.1), xlab="age",
    ylab="YoY Mortality Improvement Rate", cex.axis=1.5,cex.lab=1.5,cex=2)
  lines(agesForecast,m1backgp, col=4, lwd=2, lty=3)
  lines(c(0,100),c(0,0))

  # MP-2015 rates
  mpRate <- dplyr::filter(mp2015, gender=="male", age %in% agesForecast, year == yr)$improvement
  lines(agesForecast, mpRate, col=3, lwd=2, lty=2)

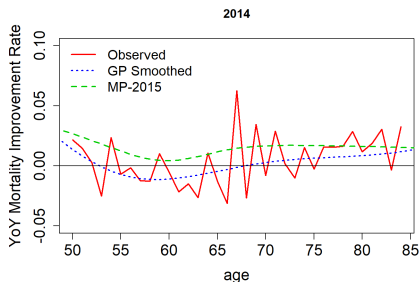
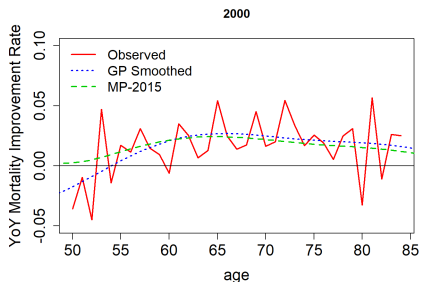
  legend("topleft",c("Observed","GP Smoothed","MP-2015"), col=c(2,4,3),lwd=rep(2,3),
    lty=c(1,3,2),cex=1.3, bty="n")
}

```



# Comparing Mortality Improvement Methods

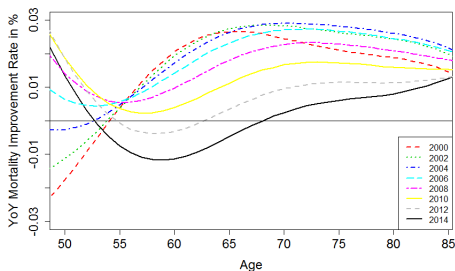
- Raw improvements extremely noisy (unsurprising)
- Smoothed methods both follow data well
- GP implies a stronger decline
  - ▶ Additional data suggests mortality deceleration



# GP Improvement Over Time

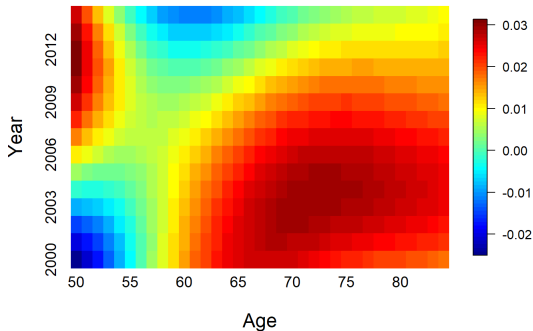
GP Improvements from 2000–2014 (in 2 year increments)

- Shape changes (flips) over time
- Consistent with MP-2015
- Generally decelerating after age 55



# GP Mortality Improvement Heatmap

- Heatmap indicates possible cohort type relation with mortality improvement



# Backward Difference & Derivatives

$f_*$  denotes the fitted GP

$$1 - \left( \frac{\exp(f_*(x_{ag}, yr))}{\exp(f_*(x_{ag}, yr - h))} \right)^{1/h} \approx - \frac{f_*(x_{ag}, yr) - f_*(x_{ag}, yr - h)}{h} \quad (3)$$

- As defined, the typical **annual mortality improvements** are **backward differences** with  $h = 1$
- Right side remains a GP by linearity
- Taking limit as  $h \rightarrow 0$  yields derivative
  - ▶ Exists (depending on covariance kernel)
- Closed form expressions for distribution of  $\frac{\partial f_*}{\partial x_{yr}}$

# GP Derivative

## Proposition

For the Gaussian Process  $f_*$  with a twice differentiable covariance kernel  $C$ , the limiting random variables

$$\frac{\partial f_*}{\partial x_{yr}}(x_{ag}, yr) \doteq \lim_{h \rightarrow 0} \frac{f_*(x_{ag}, yr + h) - f_*(x_{ag}, yr)}{h} \quad (4)$$

exist in mean square and form a Gaussian process  $\frac{\partial f_*}{\partial x_{yr}} \sim GP(m_{diff}, s_{diff})$ . Given the training set  $\mathcal{D} = (\mathbf{x}, \mathbf{y})$ , the posterior distribution of  $\frac{\partial f_*}{\partial x_{yr}}(x_*)$  has mean and variance

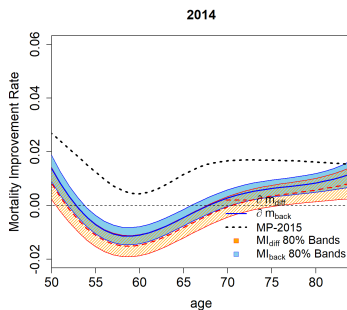
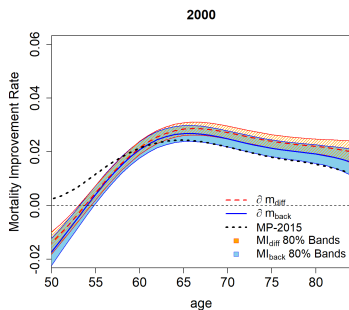
$$\begin{cases} m_{diff}(x_*) = \mathbb{E} \left[ \frac{\partial f_*}{\partial x_{yr}}(x_*) \mid \mathbf{x}, \mathbf{y} \right] = \frac{\partial C}{\partial x'_{yr}}(\mathbf{x}, x_*) (\mathbf{C} + \Sigma)^{-1} \mathbf{y}, \\ s_{diff}^2(x_*) = \text{Var} \left( \frac{\partial f_*}{\partial x_{yr}}(x_*) \mid \mathbf{x}, \mathbf{y} \right) = \frac{\partial^2 C}{\partial x_{yr} \partial x'_{yr}}(x_*, x_*) - \frac{\partial C}{\partial x'_{yr}}(\mathbf{x}, x_*) (\mathbf{C} + \Sigma)^{-1} \frac{\partial C}{\partial x_{yr}}(x_*, \mathbf{x}), \end{cases}$$

where  $\frac{\partial C}{\partial x'_{yr}}(\mathbf{x}, x_*) = \left[ \frac{\partial C}{\partial x'_{yr}}(x^1, x_*), \dots, \frac{\partial C}{\partial x'_{yr}}(x^N, x_*) \right]$  and each component is computed as the partial derivative of  $C(x, x')$ .

See Theorem 2.2.2 in Adler (2010) for more details/proof.

# Comparing Other Methods with GP Derivative

- Blue is **backwards mortality difference** (as before); red is **GP derivative**; black is MP-2015
- Analysis of other years shows deceleration begins around 2010
  - ▶ Implies mortality evolution is convex
    - ★ Justifies accelerating divergence between yearly difference and derivative methods
  - ▶ MP-2014 and MP-2015 begin to diverge around 2010
    - ★ Suggests that later years are crucial to mortality forecasts



## R Notebook Comments

- R notebook approach illustrates ease of use of GP's
- Practitioners have many options:
  - ▶ can simply change `.csv` file to their own data
  - ▶ can change output ranges for plots (e.g. plot 2016 instead of 2015)
  - ▶ have access to each plot and piece of code so programmers can specialize if needed
- Applied to US Females, Japan Male/Female, UK Male/Female data
  - ▶ Showed plausible forecasts for mortality and mortality improvement

# Conclusions

- GP's provide a variety of benefits to modeling mortality and mortality improvement
  - ▶ Bayesian approach (data driven)
  - ▶ Posterior distribution for any location
    - ★ Including distribution of mortality improvement (both yearly difference and instantaneous)
    - ★ Credible bands (historic and forecasting)
- Relatively consistent results with MP-2015
  - ▶ Four years of additional data pushes GP results in the direction that MP-2015 took compared to MP-2014 (and MP-2016, 2017 found later)
  - ▶ Differences in results is likely due to data differences than model issues
- GP framework easily handles joint analysis of mortality rates and mortality improvement



# Future Work

- Modeling annual mortality improvement directly with GP
- Monotonicity constraint:  $f \mid \frac{\partial f}{\partial f_{age}} > 0$
- Multiple populations
  - ▶ Jointly modeling male & female mortality
  - ▶ Multivariate GP of multiple countries and factors
- Modeling by cause of death

# References



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Gaussian Process Models for Mortality Rates and Improvement Factors  
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Code for Analyzing Mortality Rates and Improvements using Gaussian Processes.  
*GitHub repository* <https://github.com/jimmyrisk/GPmortalityNotebook>



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THANK YOU!